

"On certain martingales of Benoit Mandelbrot" Guest contribution (Kahane & Peyrière 1976)

♦ **Abstract.** Following his critical analysis of the random model of turbulence due to A. M. Yaglom, M 1974f{N14} and M1974c{N15} introduced his own model, which he calls "canonical." It proceeds from a brick, that is subsequently divided into $b, b^2, \dots, b^n, \dots$ similar bricks; each brick of the n -th stage is divided into b equal bricks in the $(n + 1)$ -th stage. Also given is a sequence of random variables W_p , which are independent, identically distributed, positive, have mean 1 and are indexed by the bricks P under consideration. Starting from the Lebesgue measure μ_0 on the initial brick, one constructs the sequence of measures μ_n by successive stages. Thus, μ_n has a constant density on each brick P of the n -th stage, and the density of μ_n on P is the product of W_p and the density μ_{n-1} on P . The sequence of measures μ_n is a vector martingale, and it converges towards a random measure μ . M 1974c gives results and raises problems concerning the measure μ : non-degeneracy, the moments of $\|\mu\|$, the Borel sets supporting μ and their Hausdorff dimension. Some of the conjectures of Mandelbrot have been solved by Kahane 1974 or by Peyrière 1974. Here we present these results in a refined form. Theorems 1, 2 and 3 below are due to J.-P. Kahane, Theorem 4 is due to J. Peyrière. ♦

1. Introduction, definitions, main results and history

It will be convenient to take as the initial brick the interval $[0, 1]$. The "bricks" P are then the b -adic intervals for $n = 1, 2, \dots$ and $j_k = 0, \dots, b - 1$, namely

$$I(j_1, j_2, \dots, j_n) = \left[\sum_1^n j_k b^{-k}, \sum_1^n j_k b^{-k} + b^{-n} \right].$$

Given an integer $b \geq 2$, and a positive random variable (r.v.) W with $E(W) = 1$, one denotes by $W(j_1, j_2, \dots, j_n)$ a sequence of independent r.v.'s, having the same distribution as W , and one denotes by μ_n the measure defined on $[0, 1]$, whose density on the interval $I(j_1, j_2, \dots, j_n)$ is given by $W(j_1)W(j_1, j_2) \dots W(j_1, j_2, \dots, j_n)$. Let

$$Y_n = \|\mu_n\| = b^{-n} \sum_{j_1, j_2, \dots, j_n} W(j_1)W(j_1, j_2) \dots W(j_1, j_2, \dots, j_n). \tag{1}$$

This is a nonnegative martingale, with $E(Y_n) = 1$. Hence, it converges almost surely (a.s.) towards a r.v. Y_∞ such that $E(Y_\infty) \leq 1$. In the same fashion, for all b -adic intervals I , $\mu_n(I)$ is a martingale with expectation $|I|$ which converges a.s. to a limit $\mu(I)$. Hence μ_n tends weakly a.s. to a measure μ of total mass Y_∞ .

It is convenient to write (1) in the form

$$Y_n = b^{-1} \sum_{j=0}^{b-1} W(j)Y_{n-1}(j). \tag{2}$$

The r.v. $W(j)$ and $Y_{n-1}(j)$ are mutually independent, and the $Y_{n-1}(j)$ have the same distribution as Y_{n-1} .

Consider finally the functional equation

$$Z = b^{-1} \sum_{j=0}^{b-1} W_j Z_j, \tag{3}$$

where the r.v.'s W_j and Z_j are mutually independent, the W_j having the same distribution as W , and the Z_j having the same distribution as Z . The unknown in (3) is the distribution of Z ; by an abuse of language, Z will be called solution of (3). Other solutions may exist; for example, in the case $W \equiv 1$, a Cauchy variable is a solution to (3), and it cannot be of type Y_∞ because it is neither positive nor integrable.

It will be convenient to associate with W the convex function

$$\varphi(q) = \log_b E(W^q) - (q - 1). \tag{4}$$

It is always defined for $0 \leq q \leq 1$, and can be defined for values $q > 1$. The function φ is zero at the point 1, and at most at one other point, q_{crit} . The left-side derivative of φ at the point 1 is

$$\varphi'(1-0) = E(W \log_b W) - 1 = -D \text{ (by definition of } D\text{)}.$$

We will see the role that D plays in the non-degeneracy of μ , and in the dimension of the Borel sets supporting μ . Also, we shall see the role of q_{crit} with respect to the moments of Y_∞ .

The most striking illustrations are the following. (1) $W = e^{\tau \xi - (\tau^2/2)}$, where ξ is a normal variable (this is the origin of theory); then φ is a polynomial of degree 2. (2) W has only two possible values, one of which is zero; then φ is a linear function, and $b^n Y_n$ may be interpreted as the population at time n in a birth-and-death process in which each individual gives birth to b descendants, whose probability of survival is $P(W \neq 0)$.

All these notions were introduced in M 1974f{N15} and M 1974c{N16}.

We will establish the following results.

Theorem 1. *The following statements provide equivalent conditions of non-degeneracy:*

- (α) $E(Y_\infty) = 1$,
- (β) $E(Y_\infty) > 0$,
- (γ) equation (3) has a solution Z such that $E(Z) = 1$,
- (δ) $E(W \log W) < \log b$.

Theorem 2. (Condition for the existence of finite moments). *Let $q > 1$. One has $0 < E(Y_\infty^q) < \infty$ if and only if $E(W^q) < b^{q-1}$.*

Theorem 3. (Case where Y_∞ has moments of every order). (1) *The following statements are equivalent: (α_1) $0 < E(Y_\infty^q) < \infty$ for all $q > 1$; (β_1) $\|W\|_\infty = \text{ess. sup adjst}(u \ 2)W \leq b$ and $P(W = b) < 1/b$ (strict inequality).*

(2) If (β_1) holds, one has

$$\lim_{q \rightarrow \infty} \frac{\log E(Y_\infty^q)}{q \log q} = \log_b \|W\|_\infty. \quad (5)$$

Theorem 4. (Study of the measure μ). Suppose $E(Y_\infty \log Y_\infty) < \infty$. For each $x \in [0, 1[$, denote by $I_n(x)$ the b -adic interval of order n that contains x ; its Lebesgue measure is $m(I_n(x)) = b^{-n}$. One has

$$\mu\text{-almost everywhere, } \lim_{n \rightarrow \infty} \frac{\log \mu(I_n(x))}{\log m(I_n(x))} = D = 1 - E(W \log_b W) \quad (6)$$

Corollary. The measure μ is a.s. supported by a Borel set of Hausdorff dimension D , while all Borel sets of Hausdorff dimension $< D$ have μ -measure zero.

Historical remarks. Condition (δ) of Theorem 1 can be written as $D > 0$. Kahane 1974 had only shown that

$$D > 0 \Rightarrow (\alpha) \Rightarrow (\beta) \Rightarrow (\gamma) \Rightarrow D \geq 0.$$

The role of D in the study of degeneracy had been guessed in M 1974c{N16}, Section 10.

Theorem 2 was conjectured in M 1974c and proved in Kahane 1974. The proof that we will give is simpler. Let us remark that the condition $E(W^q) < b^{q-1}$ can also be written as $\varphi(q) < 0$. If φ is zero for some $q_{crit} > 1$, this means $q < q_{crit}$.

Theorem 3 constitutes a critical comment on M 1974c, Proposition 10. It corresponds to $q_{crit} = \infty$. The proof will give some variants of Kahane 1974.

The Corollary of Theorem 4 confirms a conjecture of M 1974c{N16}, and improves on Peyrière 1974.

2. Proof of Theorem 1

Obviously $(\alpha) \Rightarrow (\beta) \Rightarrow (\gamma)$. Assume (γ) , and let Z be a solution of (3) such that $E(Z) = 1$. There exists a sequence of independent r.v.'s $W(j_1, j_2, \dots, j_n)$ ($n = 1, 2, \dots; j_k = 0, 1, \dots, b - 1$), having the same distribution as W , and a sequence of r.v.'s $Z(j_1, j_2, \dots, j_n)$ with the same distribution as Z and independent of the $W(i_1, i_2, \dots, i_n)$ for $k \leq n$, such that for all n

$$Z = b^{-n} \sum_{j_1, \dots, j_n} W(j_1)W(j_1, j_2) \dots W(j_1, j_2, \dots, j_n)Z(j_1, j_2, \dots, j_n). \quad (7)$$

Indeed, (7) reduces to (3) for $n = 1$ ($W(j) = W_j$ and $Z(j) = Z_j$), and equation (3), if applied to $Z(j_1, j_2, \dots, j_n)$, gives

$$Z(j_1, j_2, \dots, j_n) = b^{-1} \sum_{j_{n+1}} W(j_1, j_2, \dots, j_{n+1}) Z(j_1, j_2, \dots, j_n, j_{n+1}),$$

with the required conditions for the r.v.'s on the right hand side. The conditional expectation of Z with respect to the σ -field generated by the $W(j_1, \dots, j_k)$ ($k \leq n$) is Y_n as defined by (1). It follows that the martingale Y_n is uniformly integrable and that a.s. $Z = Y_\infty$ (see for example Meyer 1966, Section V8). Hence $(\gamma) \Rightarrow (\alpha)$, and furthermore (γ) implies $Z \geq 0$ a.s..

Assume again (γ) , and consequently $Z \geq 0$. For $0 < q < 1$, the function x^q is sub-additive, hence (3) yields

$$E(b^q Z^q) \leq \sum_{j=0}^{b-1} E[(W_j Z_j)^q] = bE(W^q)E(Z^q), \tag{8}$$

with $0 < E(Z^q) \leq 1$. Therefore, the function $\varphi(q)$ as defined by (4) is nonnegative on $[0, 1]$, from which it follows that $\varphi'(1 - 0) \leq 0$, that is $D \geq 0$. To go further, (8) needs to be improved.

Lemma A. $(x + y)^q \leq x^q + qy^q$ for $x \geq y > 0$, and $0 < q < 1$.

Proof. Use the fact that $y = 1$ and the formula of finite increments.

Lemma B. Let X be a positive integrable r.v., and X' a r.v. with the same distribution as X and independent of X . There exists a number $\epsilon_X > 0$ such that

$$E(X'^q 1_{X' \geq X}) \geq \epsilon_X E(X^q) \quad \text{for } 0 \leq q \leq 1.$$

Proof. Each of these expectations is a continuous function of h , and is strictly positive on $[0, 1]$.

Since the function x^q is sub-additive, it follows from (3)

$$b^q Z^q \leq \sum_{j=0}^{b-1} W_j^q Z_j^q \quad \text{a.s.}$$

From Lemma A,

$$b^q Z^q \leq qW_0^q Z_0^q + \sum_{j=1}^{b-1} W_j^q Z_j^q \quad \text{if } W_1 Z_1 \geq W_0 Z_0,$$

hence

$$E(b^q Z^q) = \sum_{j=0}^{b-1} E(W_j^q Z_j^q) - (1-q)E(W_0^q Z_0^q 1_{W_1 Z_1 \geq W_0 Z_0}),$$

which gives, using Lemma B,

$$E(b^q Z^q) \leq bE(W^q)E(Z^q) - (1-q)\varepsilon_{WZ}E(W^q)E(Z^q). \tag{9}$$

(9) is the desired refinement of (8). Dividing by $E(Z^q)$ and taking logarithms, and writing $\varepsilon = \varepsilon_{WZ}$, one has

$$\varphi(q) + \log_b(1 - \frac{(1-q)\varepsilon}{b}) \geq 0 \quad \text{on } [0, 1],$$

from which it follows that $\varphi'(1-0) + (\varepsilon/b \log b) \leq 0$, hence $D > 0$.

We have already shown that $(\alpha) \Leftrightarrow (\beta) \Leftrightarrow (\gamma) \Rightarrow (\delta)$. We will finish the demonstration by showing that (δ) implies (β)

Lemma C. $(x + y)^q \geq x^q + y^q - 2(1-q)(xy)^{q/2}$ for $x > 0, y > 0, q_0 < q < 1$.

Proof. One verifies that the function $f(t) = e^{tq} + e^{-tq} - (e^t + e^{-t})^q$ is strictly decreasing, hence has a maximum at $t=0$, where $f(1) = 2 - 2^q$ and $f'(1) = -2\log 2 < 1$. This establishes the lemma.

Here is a corollary: one has

$$\left(\sum_1^b x_j \right)^q \geq \sum_1^b x_j^q - 2(1-q) \sum_{i < j} (x_i x_j)^{q/2} \tag{10}$$

for $x_j > 0$ ($j = 1, 2, \dots, b$) and $q_0 < q < 1$. Indeed, (10) is obtained by induction from

$$\begin{aligned} \left(\sum_1^b x_j\right)^q &\geq x_1^q + \left(\sum_2^b x_j\right)^q - 2(1-q)x_1^{q/2} \left(\sum_2^b x_j\right)^{q/2} \\ &\geq x_1^q + \left(\sum_2^b x_j\right)^q - 2(1-q) \sum_{j>1} (x_1 x_j)^{q/2} \end{aligned}$$

which results from Lemma C and the sub-additivity of the function $x^{q/2}$.

Let us return to formula (2) and rewrite it provisionally as

$$Y = b^{-1} \sum_{j=0}^{b-1} W_j X_j, \tag{11}$$

where Y, W_j, X_j stand for $Y_n, W(j), Y_{n-1}(j)$. Suppose that $q_0 < q < 1$. Apply Lemma C in the form (10) with $x_{j+1} = W_j X_j$. One obtains

$$b^q Y^q \geq \sum_{j=0}^{b-1} W_j^q X_j^q - 2(1-q) \sum_{i<j} W_i^{q/2} W_j^{q/2} X_i^{q/2} X_j^{q/2}.$$

Taking expectations yields

$$b^q E(Y^q) \geq bE(W^q)E(X^q) - b(b-1)(1-q)E^2(W^{q/2})E^2(X^{q/2}).$$

By returning to our initial notations,

$$E(Y_n^q) \geq b^{1-h} E(W^q) E(Y_{n-1}^q) - b^{1-q} (b-1)(1-q) E^2(W^{q/2}) E^2(Y_{n-1}^{q/2}).$$

Taking account of $E(Y_n^q) \leq E(Y_{n-1}^q)$ (inequality of super-martingales),

$$E(Y_n^q) [1 - b^{1-h} E(W^q)] \geq -b^{1-q} (b-1)(1-q) E^2(W^{q/2}) E^2(Y_{n-1}^{q/2})$$

hence

$$E(Y_n^q) (b^{\varphi(q)} - 1) \leq b^{1-q} (b-1)(1-q) E^2(Y_{n-1}^{q/2}),$$

and, by letting $q \rightarrow 1$, we get

$$D \log b \leq (b - 1)E^2(Y_{n-1}^{1/2}).$$

Now the r.v. $Y_n^{1/2}$ are equi-integrable, since $E(Y_n) = 1$. As they converge a.s. towards Y_∞ , one has $E(Y_\infty^{1/2}) = \lim E(Y_n^{1/2})$ (see for example Meyer 1966, II.21), so that $E(Y_\infty^{1/2}) \neq 0$. This implies (β) , which concludes the demonstration of Theorem 1.

3. Proof of Theorem 2

First, suppose that (3) has a positive solution Z such that $0 < E(Z^q) < \infty$, with given $q > 1$. Because the function x^q is super-additive, one has

$$b^q Z^q \geq \sum_{j=0}^{b-1} (W_j Z_j)^q,$$

and the strict inequality holds with with positive probability, so that

$$b^q E(Z^q) > bE(W^q)E(Z^q),$$

that is $E(W^q) < b^{q-1}$.

Conversely, suppose that $E(W^q) < b^{q-1}$, so that $\varphi(q) < 0$, and let k be an integer with $k < q \leq k + 1$. As the function $x^{q/(k+1)}$ is sub-additive, one has, for $x_j \geq 0$ ($j = 1, 2, \dots, b$),

$$\begin{aligned} (x_1 + x_2 + \dots + x_b)^q &\leq (x_1^{q/k+1} + \dots + x_b^{q/k+1})^{k+1} \\ &= x_1^q + \dots + x_b^q + \sum \gamma_{\alpha_1, \dots, \alpha_b} (x_1^{\alpha_1} \dots x_b^{\alpha_b})^{q/(k+1)}. \end{aligned}$$

In the last sum, the exponents of x_j do not exceed k , the coefficients are positive, and $\sum \gamma_{\alpha_1, \alpha_2, \dots, \alpha_b} = b^{k+1} - b$.

Reconsider formula (2) in the form (11), and remark that $E(U^{q/(k+1)}) \leq E(U)$ if $U \geq 0$ and that $\prod_j E(U^{\alpha_j}) \leq E(U)E(U^k)$ if the α_j are non-negative integers such that $\sum \alpha_j = k + 1$, and at least two of the α_j be different from 0. We obtain

$$b^q E(Y^q) \leq bE(W^q)E(X^q) + (b^{k+1} - b)[E(W^k)E(X^k)]^{q/k}.$$

Hence

$$E(Y_n^q) \leq b^{1-q} E(W^q) E(Y_{n-1}^q) + b[E(W^k) E(Y_{n-1}^k)]^{q/k}.$$

Taking into account the sub-martingale inequality $E(Y_n^q) \geq E(Y_{n-1}^q)$

$$E(Y_n^q)(1 - b^{1-h} E(W^q)) \leq b[E(W^k) E(Y_n^k)]^{q/k}.$$

Letting n go to infinity, one can see that

$$E(Y_\infty^k) < \infty \Rightarrow E(Y_\infty^q) < \infty.$$

This establishes the desired result for $1 < q \leq 2$. Now assume that $q > 2$. As the hypothesis $\varphi(q) < 0$ implies $\varphi(l) < 0$ for all integer $l \leq q$, one also has

$$E(Y_\infty^{l-1}) < \infty \Rightarrow E(Y_\infty^l) < \infty$$

for $l = 2, \dots, k$. It results from the above implications that $E(Y_\infty^q) < \infty$. This concludes the proof of Theorem 2.

4. Proof of Theorem 3

Part 1. According to Theorem 2, (α_1) implies $E(W^q) < b^{q-1}$ for all $q > 1$. This implies (β_1) . Conversely, if (β_1) holds, condition (δ) of Theorem 1 is satisfied, hence $E(Y_\infty^q) > 0$. Furthermore $E(W^q) \leq b^q$ i.e. $\varphi(q) \leq 1$ for all $q > 0$. Since $\varphi(1) = 0$ and φ is convex, this means that $\varphi(q) < 0$ for all $q > 1$, that is $E(W^q) < b^{q-1}$, or $\varphi \equiv 0$, i.e. $W \equiv 1$. Hence $E(Y_\infty^q) < \infty$ for all $q > 0$. Therefore $(\alpha_1) \Leftrightarrow (\beta_1)$.

Part 2. (β_1) . Hence (by Theorem 1) $E(Y_\infty) = 1$. On the other hand, there exists an $\varepsilon > 0$ such that $\varphi(q) < \log_b(1 - \varepsilon)$ for $q \geq 2$. That is,

$$E(W^q) \leq (1 - \varepsilon)b^{q-1} \quad (q \geq 2). \tag{12}$$

Consider formula (3), with $Z = Y_\infty$, and let h be an integer ≥ 2 . One has

$$b^q Z^q = \left(\sum_{j=0}^{b-1} W_j Z_j \right)^q,$$

from where

$$b^q E(Z^q) = bE(W^q)E(Z^q) + \sum_{\substack{q_1 + \dots + q_b = q \\ q \leq q-1}} \frac{q!}{q_1! \dots q_b!} \prod_{j=1}^b E(W^{q_j}) \prod_{j=1}^b E(Z^{q_j}). \tag{13}$$

(13) together with (12) gives

$$\varepsilon b^q E(Z^q) \leq \sum_{\substack{q_1 + \dots + q_b = q \\ q \leq q-1}} \frac{q!}{q_1! \dots q_b!} \prod E(W^{q_j}) \prod E(Z^{q_j}), \tag{14}$$

hence

$$E(Z^q) \leq \frac{1}{\varepsilon} \sum_{\substack{q_1 + \dots + q_b = q \\ q \leq q-1}} \frac{q!}{q_1! \dots q_b!} \prod E(Z^{q_j}).$$

Lemma D. For all $\alpha > 0$, one has

$$\sum_{\substack{q_1 + \dots + q_b = q \\ q \leq q-1}} (q_1! \dots q_b!)^\alpha = o((q!)^\alpha) \quad (q \rightarrow \infty).$$

The proof follows immediately for $b = 2$, and continues through recurrence over b .

$c > 0$ being given, write $A_q = \sup_{l < q} (E(Z^l) / (l!)^{1+c})^{1/l}$. Then

$$A_{q+1}^q \leq \sup \left(\frac{1}{\varepsilon} \frac{\sum (q_1! \dots q_b!)^c}{(q!)^\alpha} A_{q'}^q, A_q^q \right).$$

According to Lemma D, the sequence A_q is bounded, therefore

$$E(Z^q) \leq A(c)^q (q!)^{1+c}, \text{ with } A(c) < \infty.$$

As a consequence,

$$\lim \sup_{q \rightarrow \infty} \frac{\log E(Y_\infty^q)}{q \log q} \leq 1. \tag{15}$$

If we suppose that $\|W\|_\infty = \gamma < b$, (14) gives

$$E(Z^q) \leq \frac{1}{\varepsilon} \left(\frac{\gamma}{b} \right)^q \sum \frac{q!}{q_1! \dots q_b!} \prod E(Z^{q_i}).$$

Set $B_q = \sup_{l < q} (E(Z^l)/l!)^{1/l}$.

As the number of terms in the sum Σ does not exceed q^b , one obtains

$$B_{q+1}^q \leq \sup \left(\frac{1}{\varepsilon} \left(\frac{\gamma}{b} \right)^q q^b B_q^q, B_q^q \right)$$

which means that the sequence B_q is bounded. It results that $E(e^{tZ}) < \infty$ for all small enough $t > 0$.

Now set $e^{\chi(t)} = E(e^{tZ})$. Formula (3) then becomes

$$e^{\chi(bt)} = E^b(e^{\chi(Wt)}). \tag{16}$$

The hypothesis $\|W\|_\infty = \gamma < b$ implies $\chi(bt) \leq b\chi(\gamma t)$, hence $\chi((b/\gamma)^n) = O(b^n)$ for $n \rightarrow \infty$. Setting $(b/\gamma)^K = b$, one has

$$\chi(t) = O(t^K) \quad (t \rightarrow \infty). \tag{17}$$

It is an easy exercise to show that (17) is equivalent to the existence of a positive real B such that

$$E(Z^q) \leq B^q (q!)^{1-(1/K)}.$$

However, $1 - (1/K) = \log_b \gamma$. Hence one has

$$\limsup_{q \rightarrow \infty} \frac{\log E(Y^q)}{q \log q} \leq \log_b \gamma. \tag{18}$$

Let us now choose $1 < \gamma_1 < \|W\|_\infty$ (the case $\|W\|_\infty = 1$ is straightforward). There exists an $\varepsilon > 0$ such that $E(W^q) \geq \varepsilon \gamma_1^q$. Reexamine formula (13). Since

$$\sum_{\substack{q_1 + \dots + q_b = q \\ q_i \leq q-1}} \frac{q!}{q_1! \dots q_b!} = b^q - b,$$

one has

$$\begin{aligned}
 E(Z^q) &\geq \frac{b^q - b}{b^q} \inf \prod_{j=1}^b E(W^{q_j}) E(Z^{q_j}) \\
 &\geq \left| \frac{1}{2} \right| \varepsilon^b \gamma_1^q \inf \prod_{j=1}^b E(Z^{q_j})
 \end{aligned}$$

the infimum being taken over all b -tuples (q_1, q_2, \dots, q_b) such that $q_1 + q_2 + \dots + q_b = q$ and $\sup_j q_j \leq q - 1$. Let q be a multiple of b .

The infimum is then $E^b(Z^q/b)$. Hence

$$\frac{\log E(Z^{bq})}{bq} \geq \frac{\log E(Z^q)}{q} + \log \gamma_1 + O\left(\frac{1}{q}\right),$$

with the consequence

$$\log E(Z^q) \geq \eta q \log q + O(q), \quad \eta = \log_b \gamma_1 \tag{19}$$

from where

$$\liminf_{q \rightarrow \infty} \frac{\log E(Y_\infty^q)}{q \log q} \geq \log_b \gamma_1 \tag{20}$$

(15), (18) and (20) lead to

$$\lim_{q \rightarrow \infty} \frac{\log E(Y_\infty^q)}{q \log q} = \log_b \|W\|_\infty.$$

This terminates the proof of Theorem 3.

Remark. In the case $0 < P(W = b) < 1/b$, one has $\gamma_1 = b$ in (19), from which follows that $E(e^{tZ}) = \infty$ for large enough $t > 0$.

5. Proof of Theorem 4

Let Ω be the space on which the random variables $W(j_1, j_2, \dots, j_n)$ are defined. On the product space $\Omega \times [0, 1]$, consider the probability Q defined by

$$Q(A) = E\left(\int 1_A d\mu\right).$$

Let $X_n = \sum_{j_1, \dots, j_n} W(j_1, \dots, j_n) 1_{I(j_1, \dots, j_n)}$. One has then

$$\mu_n = b^{-n} \prod_{1 \leq j \leq n} X_n$$

(with an evident abuse of notation).

Write $\mu = \mu_n \nu_n$; here ν_n is a measure whose restriction to each interval of the n -th stage is defined in an analogous fashion as for μ .

Observe that the variables $\nu_n(I(j_1, \dots, j_n))$ have the same distribution as Y_∞ and that, for fixed n , they are mutually independent. Moreover, the variables $\nu_n(I(j_1, \dots, j_n))$ and $W(k_1, \dots, k_n)$ are independent as long as the interval $I(k_1, \dots, k_p)$ is not strictly contained in the interval $I(j_1, \dots, j_n)$.

It is convenient to consider the random function

$$T_n = \sum_{j_1, \dots, j_n} b^n \nu_n(I(j_1, \dots, j_n)) 1_{I(j_1, \dots, j_n)}.$$

If u is a function defined on $[0, 1]$, constant on the intervals of the n -th stage, one has

$$\int u d\mu = \int_0^1 u(x) \mu_n(x) T_n(x) dx. \tag{21}$$

The theorem results from the two following Lemmas.

Lemma E. *If $E(W \log_b W) < 1$, then almost surely μ -almost everywhere $(1/n) \log \mu_n$ tends towards $E(W \log W)$ when $n \rightarrow +\infty$.*

Proof. We will show that

$$\sum_{n \geq 1} Q(\{X_n > e^{n-1}\}) < \infty \tag{22}$$

holds and that the series

$$\sum_{n \geq 1} \frac{1}{n} \{ \log \inf (X_n, e^{n-1}) - E[W \log(\inf (W, e^{n-1}))] \} \tag{23}$$

converges Q -almost surely. Therefore, one will have Q -almost surely $X_n \leq e^{j-1}$ beyond a certain rank and

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{j=1}^n \log \inf(X_j, e^{n-1}) = E(W \log W),$$

from which the lemma follows.

Let us begin by evaluating $Q(\{X_n > e^{n-1}\})$. One has

$$Q(\{X_n > e^{n-1}\}) = E \left[\int 1_{\{X_n > e^{n-1}\}} d\mu \right].$$

Taking account of (21) and the properties of independence of variables, we obtain

$$\begin{aligned} Q(\{X_n > e^{n-1}\}) &= E \int_0^1 1_{\{X_n(x) > e^{n-1}\}} \mu_n(x) T_n(x) dx \\ &= \int_0^1 E(X_n(x) 1_{\{X_n(x) > e^{n-1}\}}) E(\mu_{n-1}(x)) E(T_n(x)) dx \\ &= E(W 1_{\{W > e^{n-1}\}}), \end{aligned}$$

from where we have

$$\sum_{n \geq 1} Q(\{X_n > e^{n-1}\}) \leq E \left\{ W \sum_{n \geq 1} 1_{\{W > e^{n-1}\}} \right\} \leq E(W(1 + \log^+ W)),$$

which proves (22).

For ease of writing, set $X'_n = \log \inf(X_n, e^{n-1})$. We will calculate $E_Q(X'_n | X_1, \dots, X_{n-1})$. Let u be a bounded Borel function from \mathbb{R}^{n-1} into \mathbb{R} . One has

$$\begin{aligned} & \int u(X_1, \dots, X_{n-1})X'_n dQ \\ &= E \int_0^1 u(X_1(x), \dots, X_{n-1}(x))X'_n(x)\mu_n(x)T_n(x)dx \\ &= \int_0^1 E[u(X_1(x), \dots, X_{n-1}(x))\mu_{n-1}(x)]E[X'_n(x)X_n(x)]dx \\ &= E[W \log \inf(W, e^{n-1})] \int_0^1 E[u(X_1(x), \dots, X_{n-1}(x))\mu_{n-1}(x)T_{n-1}(x)]dx. \end{aligned}$$

This proves that $E_Q(X'_n | X_1, \dots, X_{n-1}) = E[W \log \inf(W, e^{n-1})]$. Therefore

$$\int (X'_n)^2 dQ = E \int_0^1 (X'_n(x))^2 \mu_n(x) T_n(x) dx = E[W(\log \inf(W, e^{n-1}))^2],$$

hence

$$\begin{aligned} & \sum_{n \geq 2} \frac{1}{n^2} \int (X'_n)^2 dQ \\ &= E \left[W \sum_{n \geq 2} \frac{1}{n^2} (\log \inf(W, e^{n-1}))^2 \right] \\ &\leq E \left[W(\log W)^2 \sum_{n \geq \sup(2, 1 + \log W)} n^{-2} + W \sum_{2 \leq n < 1 + \log W} \left(\frac{n-1}{n} \right)^2 \right] \\ &\leq E \left[W \left(\log^+ W + \frac{(\log W)^2}{\sup(1, \log W)} \right) \right]. \end{aligned}$$

The theorem on the convergence of L^2 -martingales gives (23).

Lemma F. *Suppose that $E(W \log_b W) < 1$ and that $E(Y_\infty \log Y_\infty) < \infty$. Then almost surely μ -almost everywhere $(1/n) \log v_n(I_n(x))$ tends towards $-\log b$.*

Proof. One has

$$\int T_n^{-1/2} dq = E \int_0^1 \mu_n(x) (T_n(x))^{1/2} dx = E(Y_\infty^{1/2}),$$

from where

$$\int \left(\sum_{n \geq 1} \frac{1}{n^2} T_n^{-1/2} \right) dQ < \infty.$$

As a consequence, Q -almost surely, $T_n^{-1/2} \leq n^2$ holds above a certain n , hence $\liminf_{n \rightarrow \infty} (1/n) \log T_n \geq 0$. Up to now, we have not used the second hypothesis. Now let us show that, Q almost surely, one has $\limsup \text{adjust}(u 2)_{n \rightarrow \infty} (1/n) \log T_n \leq 0$. Take a number $\alpha > 1$. Using (21),

$$E \int 1_{\{T_n > \alpha^n\}} d\mu = E(Y_\infty 1_{\{Y_\infty > \alpha^n\}}).$$

Then

$$\sum_{n \geq 1} Q(\{T_n > \alpha^n\}) = E \left(Y_\infty \sum_{n \geq 1} 1_{\{Y_\infty > \alpha^n\}} \right) \leq E(Y_\infty \log_\alpha^+ Y_\infty)$$

follows, proving that, for all $\alpha > 1$, it is Q -almost sure that $\limsup \text{adjust}(u 2)_{n \rightarrow \infty} (1/n) \log T_n \leq \log \alpha$.

The desired result follows. As $v_n(I_n(x)) = b^{-n} T_n(x)$, the lemma is proven. To prove the corollary, use Billingsley 1967, p. 136-145.

Remark. Under the only hypothesis $E(W \log_b W) < 1$, one obtains that almost surely every Borel set of dimension $< D$ is of vanishing μ -measure.

&&&&& POST-PUBLICATION APPENDIX &&&&&

AN ELEMENTARY EXPLICIT SOLUTION OF MANDELBROT'S FUNCTIONAL EQUATION $bY = \sum_{j=1}^b W_j Y_j$, BY J. PEYRIÈRE (1998)

Mandelbrot's equation starts with an integer base $b > 1$ and a random variable $W \geq 0$ satisfying $EW = 1$ and $E(W \log W) < \log b$. The unknown is a random variable Y satisfying $EY = 1$. This appendix proposes to show that an elementary explicit solution exists in the special case when $0 < W < b$ and $\Pr\{W < w\} = (w/b)^\alpha$ for $w < b$, with $\alpha = 1/(b - 1)$. In that case, the gen-

